

# 國立政治大學 110 學年度第 1 學期博士班資格考考試命題紙

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考試科目	理論統計	卷別	第一卷	考試日期	110年9月6日 星期一
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## 數理統計、機率論 (總分：50)

1. (30 pts) Suppose that we have a sample of IID observations  $X_1, \dots, X_n$  and the  $X_i$ 's have a common Lebesgue density function  $f$ . Suppose that  $f$  is twice differentiable on  $(-\infty, \infty)$  and  $f''$  is continuous and bounded on  $(-\infty, \infty)$ . For  $x_0 \in (-\infty, \infty)$ , consider the following estimator of  $f(x_0)$ :

$$\hat{f}_n(x_0) = \frac{1}{nh_n} \sum_{i=1}^n k\left(\frac{x_0 - X_i}{h_n}\right),$$

where  $\{h_n\}_{n=1}^{\infty}$  is a sequence of real numbers such that  $\lim_{n \rightarrow \infty} h_n = 0$  and

$$k(x) = \begin{cases} 1 - |x| & \text{if } |x| \leq 1; \\ 0 & \text{otherwise.} \end{cases}$$

- (a) (10 pts) Find

$$\lim_{n \rightarrow \infty} \frac{E(\hat{f}_n(x_0)) - f(x_0)}{h_n^2}.$$

- (b) (12 pts) Specify condition(s) on  $\{h_n\}_{n=1}^{\infty}$  to guarantee that for  $x_0 \in (-\infty, \infty)$ ,  $\hat{f}_n(x_0)$  converges to  $f(x_0)$  in probability as  $n \rightarrow \infty$ . Justify your answer.
- (c) (8 pts) Show that

$$\frac{\hat{f}_n(x_0) - E(\hat{f}_n(x_0))}{\sqrt{\text{Var}(\hat{f}_n(x_0))}}$$

converges to  $N(0, 1)$  in distribution as  $n \rightarrow \infty$ . Specify additional condition(s) if necessary.

2. (10 pts) Suppose that  $T, T_j: j = 1, 2, \dots$  are decision rules for the same decision problem. Suppose that for  $j \geq 1$ ,  $T_j$  is a Bayes rule with Bayes risk  $r_j$ . Suppose that the risk function of  $T$  is a constant  $r \in (-\infty, \infty)$  and

$$r = \lim_{j \rightarrow \infty} r_j.$$

Show that  $T$  is a minimax decision rule.

3. (10 pts) Suppose that we have a sample of IID observations  $X_1, \dots, X_n$ , where the distribution of  $X_1$  is  $N(\mu, 1)$ : the normal distribution of mean  $\mu$  and variance 1, and  $\mu \in (-\infty, \infty)$  is an unknown parameter. Consider the problem of estimating  $\mu$  under a loss function  $L$ , where  $L$  is defined so that the loss for estimating  $\mu$  using  $a$  when  $\mu = \mu_0$  is

$$L(a, \mu_0) = (a - \mu_0)^2$$

for  $a \in (-\infty, \infty)$  and  $\mu_0 \in (-\infty, \infty)$ .

- (a) (5 pts) Find the Bayes estimator for  $\mu$  under the loss  $L$  when the prior distribution for  $\mu$  is  $N(0, \tau^2)$ : the normal distribution of mean zero and variance  $\tau^2$ , where  $\tau$  is a positive constant.
- (b) (5 pts) Find a minimax estimator for  $\mu$ .

考試科目	理論統計	卷別	第二卷	考試日期	110年9月6日 星期一
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線性模式 (總分：50)

NOTATION:

- (i)  $\mathbf{0}_n$ : the zero  $n$ -vectors in  $\mathcal{R}^n$ .
- (ii)  $I_n$ : the identity matrix of size  $n$ .
- (iii)  $P_V$ : the projection matrix on a vector space  $V$ .
- (iv)  $\mathcal{L}(X)$ : the subspace spanned by the column vectors of a matrix  $X$ .

1. For the linear model  $\mathbf{Y} = X\boldsymbol{\beta} + \boldsymbol{\epsilon}$ , where  $X$  is an  $n \times p$  matrix and  $\boldsymbol{\epsilon}$  has mean 0 and covariance matrix  $\sigma^2 I_n$ , we say  $\eta = \mathbf{c}'\boldsymbol{\beta}$ ,  $\mathbf{c} \in \mathcal{R}^p$ , is estimable if and only if there exists a vector  $\mathbf{a}$  in  $\mathcal{R}^n$  such that  $E(\mathbf{a}'\mathbf{Y}) = \mathbf{c}'\boldsymbol{\beta}$  for all  $\boldsymbol{\beta} \in \mathcal{R}^p$ .

(a) (5pts). Please show that  $\mathbf{c}'\boldsymbol{\beta}$  is estimable if and only if  $\mathbf{c}$  lies in the row space of  $X$ . (Therefore, all components of  $\boldsymbol{\beta}$  are estimable if and only if the column vectors of  $X$  are linear independent.)

(b) (10pts). Let  $\mathbf{a}_* = P_{\mathcal{L}(X)}\mathbf{a}$ . Please show that  $\mathbf{a}'_*\mathbf{Y}$  has the smallest variance among all linear unbiased estimators of  $\eta$ . Namely,  $\mathbf{a}'_*\mathbf{Y}$  is unbiased and  $var(\mathbf{a}'_*\mathbf{Y}) \leq var(\mathbf{d}'\mathbf{Y})$  for all  $E(\mathbf{d}'\mathbf{Y}) = \eta$ .

2. (7pts). Consider the regression model  $\mathbf{Y} = X\boldsymbol{\beta} + \boldsymbol{\epsilon}$  with  $E(\boldsymbol{\epsilon}) = \mathbf{0}_n$  and  $Cov(\boldsymbol{\epsilon}) = \sigma^2 I_n$ . Let us decompose  $\mathbf{Y} = \begin{pmatrix} \mathbf{Y}_1 \\ \mathbf{Y}_2 \end{pmatrix}$  and  $X = \begin{pmatrix} X_1 \\ X_2 \end{pmatrix}$ , where  $\mathbf{Y}_i$  and  $X_i$  have  $n_i$  rows for  $i = 1, 2$  and  $n = n_1 + n_2$ . Let SSE, SSE<sub>1</sub> and SSE<sub>2</sub> denote the usual error sum of squares for regression of  $\mathbf{Y}$  on  $X$ ,  $\mathbf{Y}_1$  on  $X_1$  and  $\mathbf{Y}_2$  on  $X_2$  respectively. Namely,  $SSE = \|\mathbf{Y} - P_{\mathcal{L}(X)}\mathbf{Y}\|^2$ ,  $SSE_1 = \|\mathbf{Y}_1 - P_{\mathcal{L}(X_1)}\mathbf{Y}_1\|^2$ , and  $SSE_2 = \|\mathbf{Y}_2 - P_{\mathcal{L}(X_2)}\mathbf{Y}_2\|^2$ . Which one of the following statements is correct? Justify your answer.

- (i)  $SSE \leq SSE_1 + SSE_2$ ; (ii)  $SSE = SSE_1 + SSE_2$ ; (iii)  $SSE \geq SSE_1 + SSE_2$ ;
- (iv) none of the above.

3. Here we consider two linear regression models as below.

$$\text{Model (A): } \mathbf{Y} = X_1\boldsymbol{\beta}_1 + \boldsymbol{\epsilon};$$

$$\text{Model (B): } \mathbf{Y} = X_1\boldsymbol{\beta}_1 + X_2\boldsymbol{\beta}_2 + \boldsymbol{\delta}.$$

Here  $X_1$  and  $X_2$  are  $n \times p_1$  and  $n \times p_2$  constant matrices,  $\boldsymbol{\beta}_1$  and  $\boldsymbol{\beta}_2$  are unknown  $p_1$ -vector and  $p_2$ -vector, and  $\boldsymbol{\epsilon}$ ,  $\boldsymbol{\delta}$  are random vectors with mean  $\mathbf{0}_n$  and covariance matrix  $\sigma^2 I_n$ . Suppose matrix  $X_{n \times (p_1+p_2)} = [X_1 \ X_2]$  has full column rank. Denote  $\hat{\boldsymbol{\beta}}_1$  the unique LSE under model (A) and  $\tilde{\boldsymbol{\beta}}_1$  and  $\tilde{\boldsymbol{\beta}}_2$  the unique LSE under model (B). Also let  $\hat{\sigma}^2 = \frac{\|\mathbf{Y} - X_1\hat{\boldsymbol{\beta}}_1\|^2}{n-p_1}$  and  $\tilde{\sigma}^2 = \frac{\|\mathbf{Y} - X_1\tilde{\boldsymbol{\beta}}_1 - X_2\tilde{\boldsymbol{\beta}}_2\|^2}{n-p_1-p_2}$  be the usual estimates of  $\sigma^2$  based on model (A) and (B) respectively.

- (a) (8pts). If model (A) is true, it is known that  $\hat{\boldsymbol{\beta}}_1$  and  $\hat{\sigma}^2$  are both unbiased. What about  $\tilde{\boldsymbol{\beta}}_1$  and  $\tilde{\sigma}^2$ ? Justify your answer.
- (b) (10pts). If model (B) is true, we know  $\tilde{\boldsymbol{\beta}}_1$  and  $\tilde{\sigma}^2$  are unbiased. However,  $\hat{\boldsymbol{\beta}}_1$  and  $\hat{\sigma}^2$  are not. Please find bias for both of them (in terms of  $X_1, X_2, \boldsymbol{\beta}_1, \boldsymbol{\beta}_2, \sigma^2, \dots$ ).
- (c) (10pts). Suppose model (B) is true. Please find a necessary and sufficient condition (on  $X_1$  and  $X_2$ ) such that  $\hat{\boldsymbol{\beta}}_1$  is unbiased. Justify your answer. Please also make a comment on the bias of  $\hat{\sigma}^2$  under this condition.